

3. FUNCTIONAL ADDITIVE MODELING

$$E(Y - Y|k) = b_k \quad E(\zeta_m|k) = b_{km} \quad (1)$$

$$E(Y|X) = Y + \sum_{k=1}^{\infty} f_k(\cdot) \quad (2)$$

$$E(Y(t)|X) = Y(t) + \sum_{k=1}^{\infty} \sum_{m=1}^{\infty} f_{km}(\cdot) m(t), \quad (3)$$

$$E f_k(\cdot) = \beta, \quad k = 1, \dots, \quad (4)$$

$$E f_{km}(\cdot) = \beta, \quad k = 1, \dots, m = 1, \dots \quad (5)$$

$$E(Y - Y|k) = E\{E(Y - Y|X)|k\} = E\left\{\sum_{j=1}^{\infty} f_j(\cdot|k)\right\} = f_k(\cdot), \quad (6)$$

$$E(\zeta_m|k) = E\{E(\zeta_m|X)|k\} = E\left\{\sum_{j=1}^{\infty} f_{jm}(\cdot|k)\right\} = f_{km}(\cdot), \quad (7)$$

$$f_k(\cdot) = E(Y - Y|k) \quad f_{km}(\cdot) = E(\zeta_m|k) \quad k, m = 1, \dots, \quad (8)$$

4. FITTING OF FUNCTIONAL ADDITIVE MODELS

$$\hat{Y}_k = \hat{Y} + \sum_{k=1}^K \hat{f}_k(\cdot) \quad (9)$$

principal analysis by conditional expectation (E)

$$\hat{f}_k(\cdot) = \hat{f}_{km}(\cdot) \quad (10)$$

$$\{ \hat{y}_{ik}, Y_i \}_{i=1, \dots, n} \quad \hat{y}_{ik} \quad (11)$$

$$\sum_{i=1}^n K \left(\frac{\hat{y}_{ik} - x}{h_k} \right) \{ Y_i - \beta - \beta(x - \hat{y}_{ik}) \} \quad (12)$$

$$\hat{f}_k(x) = \hat{\beta}(x) - \bar{Y} \quad h_k \quad (13)$$

$$\{\hat{\zeta}_{im}\}_{i=1, \dots, n} \quad y \quad ()$$

$$\sum_{i=1}^n K \left(\frac{\hat{ik} - x}{h_{mk}} \right) \{\hat{\zeta}_{im} - \beta - \beta(x - \hat{ik})\} \quad ()$$

$$h_{mk} \quad \beta \quad \beta \quad \hat{f}_{mk}(x) = \hat{\beta}(x) \quad ()$$

$$\hat{E}(Y|X) = \bar{Y} + \sum_{k=1}^K \hat{f}_k(\cdot) \quad ()$$

$$R = - \frac{\sum_{i=1}^n \{Y_i - E(Y_i|X_i)\}}{\sum_{i=1}^n (Y_i - \bar{Y})} \quad ()$$

$$\hat{R} = - \frac{\sum_{i=1}^n \{Y_i - \hat{E}(Y_i|X_i)\}}{\sum_{i=1}^n (Y_i - \bar{Y})} \quad ()$$

$$E(Y_i|X_i) \quad \hat{E}(Y_i|X_i) \quad i \quad ()$$

$$\hat{E}\{Y(t)|X\} = \hat{Y}(t) + \sum_{m=1}^M \sum_{k=1}^K \hat{f}_{mk}(\cdot) \hat{m}(t), \quad t \in \mathcal{T}, \quad ()$$

$$R = - \frac{\sum_{i=1}^n \int [Y_i(t) - E\{Y_i(t)|X_i\}] dt}{\sum_{i=1}^n \int \{Y_i(t) - \bar{Y}(t)\} dt} \quad ()$$

$$\hat{R} = - \frac{\sum_{i=1}^n \sum_{l=1}^{m_i} [V_{il} - \hat{E}\{Y_i(t_{il})|X_i\}] (t_{il} - t_{i,l-})}{\sum_{i=1}^n \sum_{l=1}^{m_i} \{V_{il} - \bar{Y}(t_{il})\} (t_{il} - t_{i,l-})} \quad ()$$

$$\hat{E}\{Y_i(t)|X_i\} \quad i \quad t_{il} \quad E\{Y_i(t)|X_i\} \quad ()$$

5. THEORETICAL RESULTS

$$\hat{\zeta}_{im} \quad \hat{\zeta}_{im} \quad k = 1, \dots, K \quad m = 1, \dots, M \quad ()$$

$$|\hat{\zeta}_{im} - \zeta_{im}| \quad i \quad i = 1, \dots, n \quad ()$$

$$f_k \quad f_{mk} \quad () \quad ()$$

$$\{\hat{\zeta}_{im}, Y_i\} \quad \{\hat{\zeta}_{im}, \hat{\zeta}_{im}\} \quad i = 1, \dots, n \quad ()$$

$$K(n) \rightarrow \infty \quad M = M(n) \rightarrow \infty \quad n \rightarrow \infty \quad ()$$

Theorem 1. $k \geq j, j \leq k$ $() () () ()$

$$|\hat{f}_k(x) - f_k(x)| \xrightarrow{p} 0 \quad ()$$

$$|\hat{f}_{km}(x) - f_{km}(x)| \xrightarrow{p} 0 \quad ()$$

$$|\hat{f}_k(x) - f_k(x)| \quad \tilde{\vartheta}_{mk}(x) = |\hat{f}_{mk}(x) - f_{mk}(x)| \quad \tilde{\vartheta}_k(x) = \quad ()$$

Theorem 2. $() () () () () ()$

$$\hat{E}(Y|X) - E(Y|X) \xrightarrow{p} 0 \quad ()$$

$$\hat{E}(Y|X) = \bar{Y} + \sum_{k=1}^K \hat{f}_k(\cdot) \quad ()$$

$$\hat{E}\{Y(t)|X\} - E\{Y(t)|X\} \xrightarrow{p} 0 \quad ()$$

$$\hat{E}\{Y(t)|X\} = \hat{Y}(t) + \sum_{k=1}^K \sum_{m=1}^M \hat{f}_{mk}(\cdot) \hat{m}(t) \quad ()$$

$$|\hat{E}(Y|X) - E(Y|X)| \quad \vartheta_n^* = |\hat{E}(Y(t)|X) - E(Y(t)|X)| \quad ()$$

6. SIMULATION STUDIES

$$X_i \quad X(s) = s + (s) \leq s \leq \quad ()$$

$$(s) = - (s/\sqrt{\cdot})/\sqrt{\cdot} \quad (s) = (s/\sqrt{\cdot})/\sqrt{\cdot} \leq \quad ()$$

$$\epsilon_{ij} \sim (\cdot, \cdot) \quad () \quad ik \sim \mathcal{N}(\cdot, k) \quad ()$$

$$ik = \sqrt{k}(Z_{ik} - \cdot)/\cdot \quad Z_{ik} \sim (\cdot, \cdot) \quad k = \cdot, \quad ()$$

$$\hat{f}_{km}(x) = \sum_{i=1}^n (\hat{\zeta}_{im} - \bar{\zeta}_{.m}) (\hat{\zeta}_{ik} - \bar{\zeta}_{.k}) \hat{y}_k$$

$$Y_i(t) = Y(t) + \zeta_i(t) \quad (t) \leq t \leq$$

$$f_k(x) = x - k \quad f_{km}, k =$$

$$\zeta_i = \sum_{k=1}^m (\hat{\zeta}_{ik} - \bar{\zeta}_{.k}) \quad ()$$

$$\zeta_i = \sum_{l=1}^m \varepsilon_{il} \quad ()$$

$$Y_i = X_i^* Y_i^* U_{ij} V_{il} U_{ij}^* V_{il}^* \quad (E)$$

$$E_i = \frac{(Y_i^* - \hat{Y}_i^*)}{Y_i^*} \quad ()$$

$$E_{i,f} = \frac{\int (Y_i^*(t) - \hat{Y}_i^*(t)) dt}{\int Y_i^*(t) dt}, \quad ()$$

$$\hat{f}_k(x) = \sum_{i=1}^n (Y_i - \bar{Y}) (\hat{\zeta}_{ik} - \bar{\zeta}_{.k}) \hat{y}_k$$

$$\bar{Y}_i = \sum_{i=1}^n Y_i/n \quad \bar{\zeta}_{.k} = \sum_{i=1}^n \hat{\zeta}_{ik}/n$$

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	E ()	n =

T_1	T_2
y_1	y_2
y_2	y_3
y_3	y_4
y_4	y_5
y_5	y_6
y_6	y_7
y_7	y_8
y_8	y_9
y_9	y_{10}
y_{10}	y_{11}
y_{11}	y_{12}
y_{12}	y_{13}
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y_{94}	y_{95}
y_{95}	y_{96}
y_{96}	y_{97}
y_{97}	y_{98}
y_{98}	y_{99}
y_{99}	y_{100}

7. APPLICATION TO GENE EXPRESSION TIME COURSE DATA

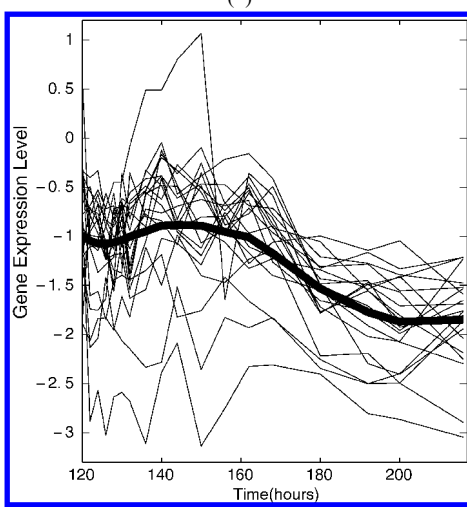
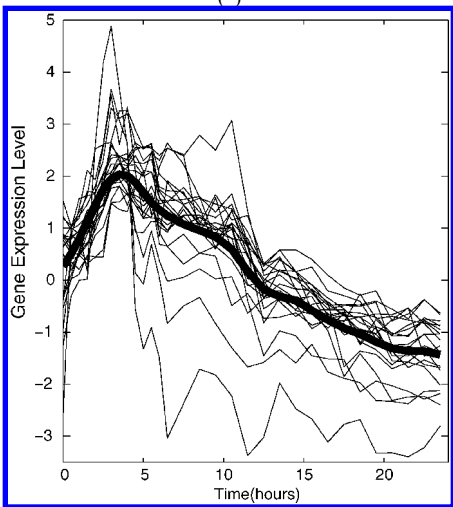
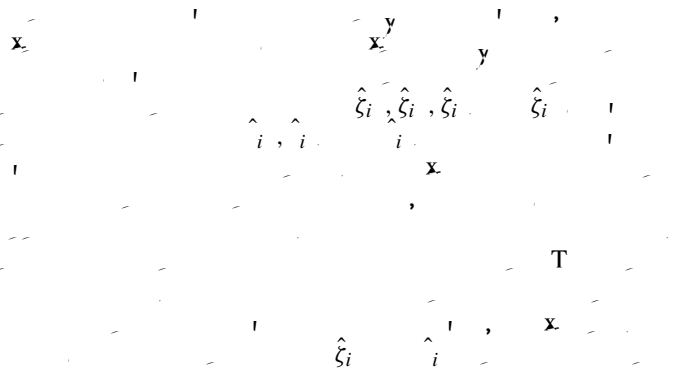
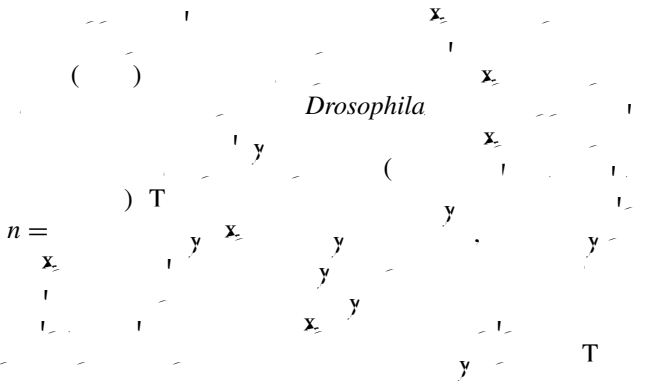
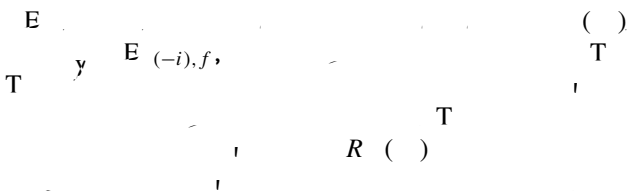
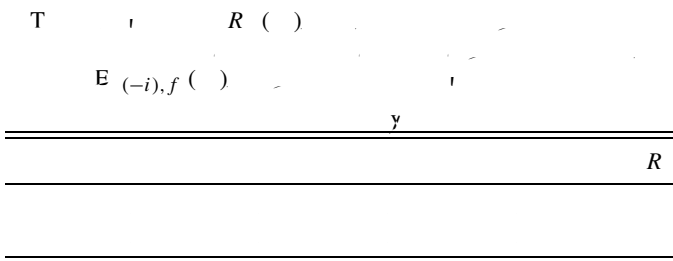
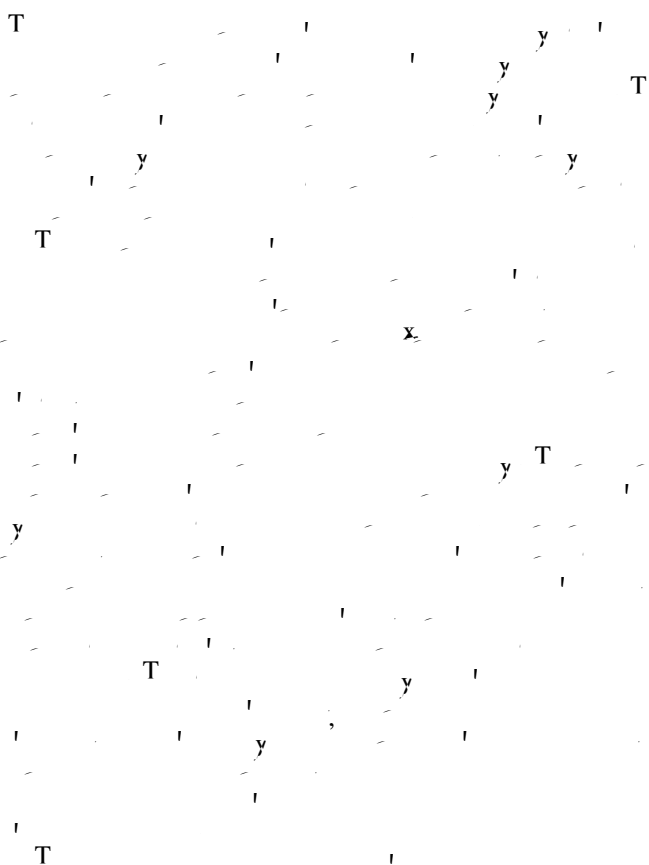
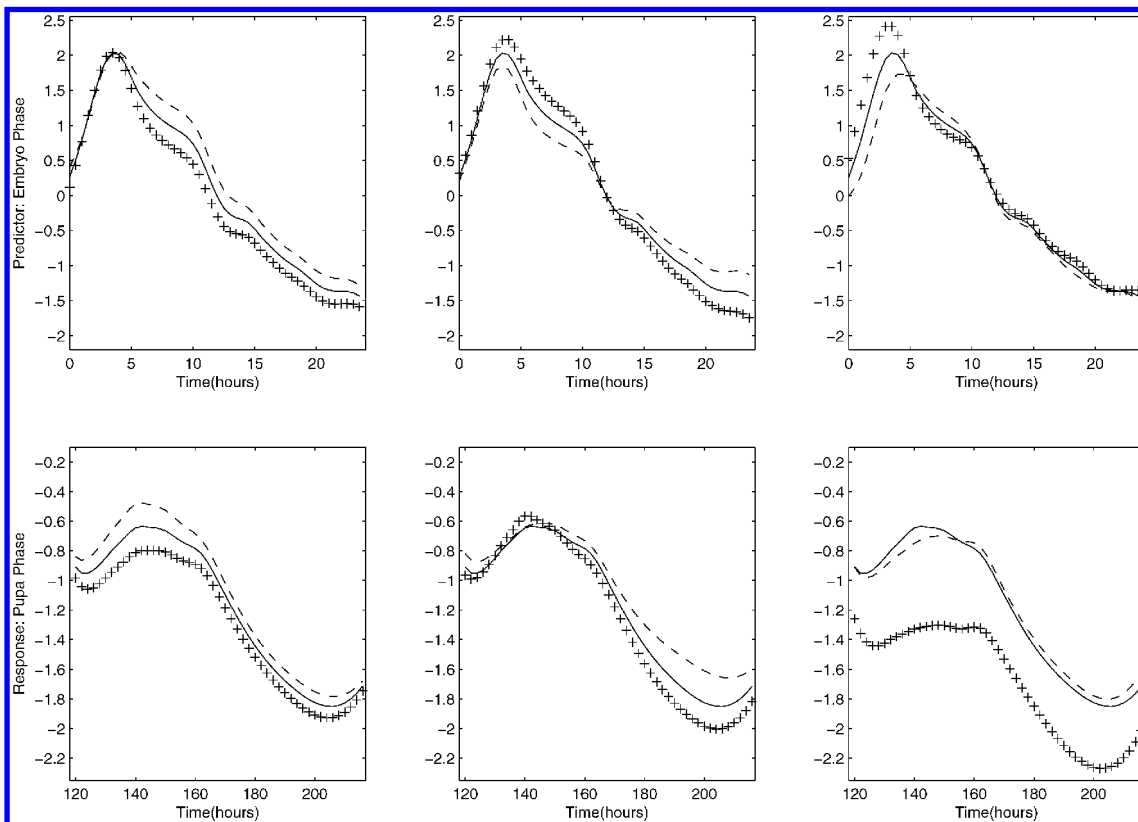
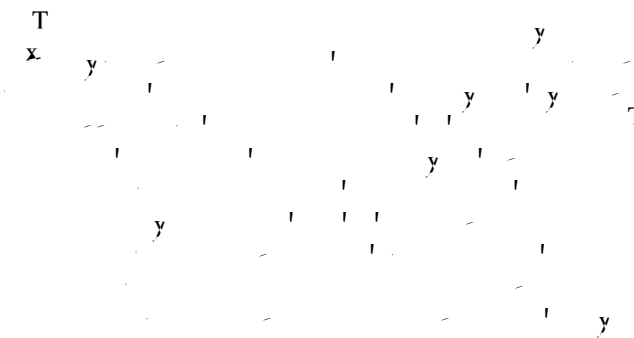


Figure 7: Application to Gene Expression Time Course Data. The figure displays two scatter plots and two line plots. The top two plots show raw data points for gene expression levels over time for Drosophila. The bottom two plots show the same data with a fitted curve overlaid, illustrating the functional additive model fit. The left plot covers the time interval from 0 to 20 hours, while the right plot covers the interval from 120 to 200 hours.



8. CONCLUDING REMARKS



$$\hat{X}(s) + \alpha \hat{k}(s) \quad \alpha = (\quad) \quad \alpha = (\quad) \quad \alpha = (\quad) \quad k = (\quad) \quad k = (\quad)$$

$$\hat{Y}(t) + \sum_{m=1}^M \{ \hat{f}_{km}(\alpha) + \sum_{\ell \neq k} \hat{f}_{\ell m}(\alpha) \} \hat{m}(t)$$

$$h_X^* = \int_{a-|S|}^b \{s - a\} ds / |S|, \quad b = \int_{a-|S|}^b \{s - a\} ds / |S|$$

$$\hat{h}_X = \int \{\hat{V}_X(s) - \tilde{G}_X(s)\} ds / |S| \quad ()$$

$$\hat{h}_X > \hat{h}_X = E \int_{k, k \geq} \{ \hat{h}_X, \hat{h}_X \} \rightarrow \mathbb{R}$$

$$\begin{aligned}
 & \mathbf{x}_k \quad t \in T \\
 & \sum_{k=1}^K [f_k''(k)|h_k + n^{-1} \{ (Y|k) \} / p_k^{-1}(k) \times \\
 & \quad h_k^{-1}] \rightarrow \\
 & \sum_{k=1}^K \sum_{m=1}^M [f_{mk}''(k) m(t)|h_{mk} + n^{-1} \{ (\zeta_m|k) \} / \times \\
 & \quad p_k^{-1}(k) | m(t)|h_{mk}^{-1}] \rightarrow
 \end{aligned}$$

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 y () T *Scandinavian Journal of Statistics*
 T x () *Canadian Journal of Statistics*

Supplement for “Functional Additive Models”

1. NOTATIONS AND AUXILIARY RESULTS

Covariance operators are denoted by \mathcal{G}_X , $\widehat{\mathcal{G}}_X$, generated by kernels G_X , \widehat{G}_X ; i.e., $\mathcal{G}_X(f) = \int_{\mathcal{T}} G_X(s, t)f(s)ds$, $\widehat{\mathcal{G}}_X(f) = \int_{\mathcal{T}} \widehat{G}_X(s, t)f(s)ds$ for any $f \in L^2(\mathcal{T})$. Define

$$\begin{aligned} D_X &= \int_{\mathcal{T}^2} \{\widehat{G}_X(s, t) - G_X(s, t)\}^2 dsdt, & \delta_k^X &= \min_{1 \leq j \leq k} (\lambda_j - \lambda_{j+1}), \\ K_0 &= \inf\{j \geq 1 : \lambda_j - \lambda_{j+1} \leq 2D_X\} - 1, & \pi_k^X &= 1/\lambda_k + 1/\delta_k^X. \end{aligned} \quad (32)$$

Let $K = K(n)$ denote the numbers of leading eigenfunctions included to approximate X as sample size n varies; i.e., $\widehat{X}_i(s) = \widehat{\mu}_X(s) + \sum_{k=1}^K \widehat{\xi}_{ik} \widehat{\phi}_k(s)$. Analogously, define the quantities \mathcal{G}_Y , $\widehat{\mathcal{G}}_Y$, D_Y , δ_m^Y , π_m^Y , M_0 and M for the process Y , for the case of functional responses. The following lemma gives the weak uniform convergence rates for the estimators of the FPCs, setting the stage for the subsequent developments. The proof is in Section 2.

Lemma 1 Under A , A , C , C and C ,

$$\sup_{t \in \mathcal{S}} |\widehat{\mu}_X(s) - \mu_X(s)| = O_p\left(\frac{1}{\sqrt{nb_X}}\right), \quad \sup_{s_1, s_2 \in \mathcal{S}} |\widehat{G}_X(s_1, s_2) - G_X(s_1, s_2)| = O_p\left(\frac{1}{\sqrt{nh_X^2}}\right), \quad (33)$$

and as a consequence, $\widehat{\sigma}_X^2 - \sigma_X^2 = O_p(n^{-1/2}h_X^{-2} + n^{-1/2}h_X^{*-1})$. Considering eigenvalues λ_k of \mathcal{G}_X of u tipicity one, $\widehat{\phi}_k$ can be chosen such that

$$P\left(\sup_{1 \leq k \leq K_0} |\widehat{\lambda}_k - \lambda_k| \leq D_X\right) = 1, \quad \sup_{s \in \mathcal{S}} |\widehat{\phi}_k(s) - \phi_k(s)| = O_p\left(\frac{\pi_k^X}{\sqrt{nh_X^2}}\right), \quad k = 1, \dots, K_0, \quad (34)$$

where D_X , π_k^X and K_0 are defined in

Analogously, under B , B , C , C and C ,

$$\sup_{t \in \mathcal{T}} |\widehat{\mu}_Y(t) - \mu_Y(t)| = O_p\left(\frac{1}{\sqrt{nb_Y}}\right), \quad \sup_{t_1, t_2 \in \mathcal{T}} |\widehat{G}_Y(t_1, t_2) - G_Y(t_1, t_2)| = O_p\left(\frac{1}{\sqrt{nh_Y^2}}\right), \quad (35)$$

and as a consequence, $\widehat{\sigma}_Y^2 - \sigma_Y^2 = O_p(n^{-1/2}h_Y^{-2} + n^{-1/2}h_Y^{*-1})$. Considering eigenvalues ρ_m of \mathcal{G}_Y of u tipicity one, $\widehat{\psi}_m$ can be chosen such that

$$P\left(\sup_{1 \leq m \leq M_0} |\widehat{\rho}_m - \rho_m| \leq D_Y\right) = 1, \quad \sup_{t \in \mathcal{T}} |\widehat{\psi}_m(t) - \psi_m(t)| = O_p\left(\frac{\pi_m^Y}{\sqrt{nh_Y^2}}\right), \quad m = 1, \dots, M_0, \quad (36)$$

here $D_{\mathbf{Y}}$, $\pi_{\mathbf{k}}^{\mathbf{Y}}$ and M_0 re de ned n ogous y to for process Y

Recall that $\|f\|_{\infty} = \sup_{\mathbf{x} \in \mathcal{A}} |f(t)|$ for an arbitrary function f with support \mathcal{A} , and $\|g\| = \sqrt{\int_{\mathcal{A}} g^2(t) dt}$ for any $g \in L^2(\mathcal{A})$ and define

$$\begin{aligned}
 \theta_{\mathbf{ik}}^{(1)} &= c_1 \|X_{\mathbf{i}}\| + c_2 \|X_{\mathbf{i}} X'_{\mathbf{i}}\|_{\infty} \mathbf{x}^* + c_3, & Z_{\mathbf{k}}^{(1)} &= \sup_{s \in \mathcal{S}} |\hat{\phi}_{\mathbf{k}}(s) - \phi_{\mathbf{k}}(s)|, \\
 \theta_{\mathbf{ik}}^{(2)} &= \mathbf{1} + \|\phi_{\mathbf{k}} \phi'_{\mathbf{k}}\|_{\infty} \mathbf{x}^*, & Z_{\mathbf{k}}^{(2)} &= \sup_{s \in \mathcal{S}} |\hat{\mu}_{\mathbf{x}}(s) - \mu_{\mathbf{x}}(s)|, \\
 \theta_{\mathbf{ik}}^{(3)} &= c_4 \|X_{\mathbf{i}}\|_{\infty} + c_5 \|X'_{\mathbf{i}}\|_{\infty} + c_6, & Z_{\mathbf{k}}^{(3)} &= \|\phi'_{\mathbf{k}}\|_{\infty} \mathbf{x}^*, \\
 \theta_{\mathbf{ik}}^{(4)} &= \left| \sum_{j=2}^{n_i} \epsilon_{ij} \phi_{\mathbf{k}}(s_{ij}) (s_{ij} - s_{i,j-1}) \right|, & Z_{\mathbf{k}}^{(4)} &\equiv \mathbf{1}, \\
 \theta_{\mathbf{ik}}^{(5)} &= \sum_{j=2}^{n_i} |\epsilon_{ij}| (s_{ij} - s_{i,j-1}), & &
 \end{aligned}$$

Lemma 2 For $\theta_{\mathbf{ik}}^{(\cdot)}$, $Z_{\mathbf{k}}^{(\cdot)}$, $\vartheta_{\mathbf{im}}^{(\cdot)}$ and $Q_{\mathbf{m}}^{(\cdot)}$ as defined in (38) and (39).

$$|\hat{\xi}_{\mathbf{ik}}^I - \xi_{\mathbf{ik}}| \leq \sum_{=1}^5 \theta_{\mathbf{ik}}^{(\cdot)} Z_{\mathbf{k}}^{(\cdot)}, \quad |\hat{\zeta}_{\mathbf{im}}^I - \zeta_{\mathbf{im}}| \leq \sum_{=1}^5 \vartheta_{\mathbf{im}}^{(\cdot)} Q_{\mathbf{m}}^{(\cdot)}. \quad (39)$$

The proof is in Section 2. In the sequel we suppress the superscript I in the FPC estimates $\hat{\xi}_{\mathbf{ik}}^I$ and $\hat{\zeta}_{\mathbf{im}}^I$.

Recall that the sequences of bandwidths $h_{\mathbf{k}}$ and $h_{\mathbf{mk}}$ are employed to obtain the estimates $\hat{f}_{\mathbf{k}}$ and $\hat{f}_{\mathbf{mk}}$ for the regression functions $f_{\mathbf{k}}$ and $f_{\mathbf{mk}}$, and that the density of $\xi_{\mathbf{k}}$ is denoted by $p_{\mathbf{k}}$. Define

$$\begin{aligned} \theta_{\mathbf{k}}(x) &= p_{\mathbf{k}}(x) \left\{ \frac{\pi_{\mathbf{k}}^{\mathbf{x}}}{\sqrt{nh_{\mathbf{x}}^2}} + \frac{1}{\sqrt{nb_{\mathbf{x}}}} + \sqrt{\pi_{\mathbf{x}}^*} \right\}, \\ \vartheta_{\mathbf{mk}}(x) &= p_{\mathbf{k}}(x) \left\{ \frac{\pi_{\mathbf{m}}^{\mathbf{y}}}{\sqrt{nh_{\mathbf{y}}^2}} + \frac{1}{\sqrt{nb_{\mathbf{y}}}} + \sqrt{\pi_{\mathbf{y}}^*} \right\}. \end{aligned} \quad (40)$$

The weak convergence rates $\tilde{\theta}_{\mathbf{k}}$ and $\tilde{\vartheta}_{\mathbf{mk}}$ of the regression function estimators $\hat{f}_{\mathbf{k}}(x)$ and $\hat{f}_{\mathbf{mk}}(x)$ (see Theorem 1) are as follows,

$$\begin{aligned} \tilde{\theta}_{\mathbf{k}}(x) &= \frac{\theta_{\mathbf{k}}(x)}{h_{\mathbf{k}}} + \frac{1}{2} |f_{\mathbf{k}}''(x)| h_{\mathbf{k}}^2 + \sqrt{\frac{\text{var}(Y|x) \|K_1\|^2}{p_{\mathbf{k}}(x) n h_{\mathbf{k}}}}, \\ \tilde{\vartheta}_{\mathbf{mk}}(x) &= \frac{\theta_{\mathbf{k}}(x)}{h_{\mathbf{mk}}} + \vartheta_{\mathbf{mk}}(x) + \frac{1}{2} |f_{\mathbf{mk}}''(x)| h_{\mathbf{mk}}^2 + \sqrt{\frac{\text{var}(\zeta_{\mathbf{m}}|x) \|K_1\|^2}{p_{\mathbf{k}}(x) n h_{\mathbf{mk}}}}. \end{aligned} \quad (41)$$

Considering the predictions $\hat{E}(Y|X)$ for the scalar response case and $\hat{E}\{Y(t)|X\}$ for the functional response case, the numbers of eigenfunctions K and M used for approximating the infinite dimensional processes X and Y generally tend to infinity as the sample size n increases. We require $K \leq K_0$ and $M \leq M_0$ in (A6). Since it follows from (35) that $K_0 \rightarrow \infty$, as long as all eigenvalues λ_j are of multiplicity 1, and analogously for M_0 , this is not a strong restriction. Denote the set of positive integers by \mathcal{N} and $\mathcal{N}_{\mathbf{k}} = \{1, \dots, k\}$. Convergence rates θ_n^* and ϑ_n^* for the predictions (20) and

(21) are as follows,

$$\begin{aligned} \theta_n^* &= \sum_{k=1}^K \left\{ \frac{\theta_k(\xi_k)}{h_k} + \frac{1}{2} |f_k''(\xi_k)| h_k^2 + \sqrt{\frac{\text{var}(Y|\xi_k) \|K_1\|^2}{p_k(\xi_k) n h_k}} \right\} + \left| \sum_{k \geq K+1} f_k(\xi_k) \right|, \quad (42) \\ \vartheta_n^* &= \sum_{k=1}^K \sum_{m=1}^M \left\{ \left(\frac{\theta_k(\xi_k)}{h_{mk}} + \vartheta_{mk}(\xi_k) \right) |\psi_m(t)| + \frac{1}{2} |f_{mk}''(\xi_k)| |\psi_m(t)| h_{mk}^2 + \sqrt{\frac{\text{var}(\zeta_m|\xi_k) \|K_1\|^2}{p_k(\xi_k) n h_k}} |\psi_m(t)| \right. \\ &\quad \left. + \frac{\pi_m^y |f_{mk}(\xi_k)|}{h_{mk}} \right\} \end{aligned}$$

Noting $\hat{\xi}_{ik} = \hat{\eta}_{ik} + \hat{\tau}_{ik}$, one finds

$$|\hat{\xi}_{ik} - \xi_{ik}| \leq \{|\hat{\eta}_{ik} - \tilde{\eta}_{ik}| + |\tilde{\eta}_{ik} - \xi_{ik}| + |\hat{\tau}_{ik}|\}. \quad (43)$$

Without loss of generality, assume $\|\phi_k\|_\infty \geq 1$, $\|\phi'_k\|_\infty \geq 1$, $\|X_i\|_\infty \geq 1$ and $\|X'_i\|_\infty \geq 1$.

For $\theta_{ik}^{(\ell)}$ and $Z_k^{(\ell)}$ (37), $\ell = 1, \dots, 5$, the first term on the r.h.s. of (43) is bounded by

$$\begin{aligned} & \left\{ \sum_{j=2}^{n_i} [|X_i(s_{ij}) - \hat{\mu}(s_{ij})| \cdot |\hat{\phi}_k(s_{ij}) - \phi_k(s_{ij})| + |\hat{\mu}(s_{ij}) - \mu(s_{ij})| \cdot |\phi_k(s_{ij})|] (s_{ij} - s_{i,j-1}) \right\} \\ & \leq \left\{ \sum_{j=1}^{n_i} [|X_i(s_{ij})| + |\mu(s_{ij})| + 1]^2 (s_{ij} - s_{i,j-1}) \right\}^{1/2} \left\{ \sum_{j=2}^{n_i} [\hat{\phi}_k(s_{ij}) - \phi_k(s_{ij})]^2 (s_{ij} - s_{i,j-1}) \right\}^{1/2} \\ & \quad + \left\{ \sum_{j=1}^{n_i} [\hat{\mu}(s_{ij}) - \mu(s_{ij})]^2 (s_{ij} - s_{i,j-1}) \right\}^{1/2} \left\{ \sum_{j=2}^{n_i} \phi_k^2(s_{ij}) (s_{ij} - s_{i,j-1}) \right\}^{1/2} \\ & \leq \theta_{ik}^{(1)} Z_k^{(1)} + \theta_{ik}^{(2)} Z_k^{(2)}. \end{aligned}$$

The second term on the r.h.s. of (43) has the upper bound

$$|\tilde{\eta}_{ij} - \xi_{ik}| \leq \|(X_i + \mu)' \phi_k + (X_i + \mu) \phi'_k\|_\infty \cdot \mathbf{x} \leq \theta_{ik}^{(3)} Z_k^{(3)}.$$

From the above, the third term on the r.h.s. of (43) is bounded by $(\theta_{ik}^{(4)} Z_k^{(4)} + \theta_{ik}^{(5)} Z_k^{(5)})$.

□

Proof of Theore For simplicity, denote " $\sum_{i=1}^n$ " by " \sum_i ", $w_i = K_1 \{(x - \xi_{ik})/h_k\}/(nh_k)$, $\hat{w}_i = K_1 \{(x - \hat{\xi}_{ik})/h_k\}/(nh_k)$, and write $\theta_k = \theta_k(x)$. From (12), the local linear estimator $\hat{f}_k(x)$ of the regression function $f_k(x)$ can be explicitly written as

$$\hat{f}_k(x) = \frac{\sum_i \hat{w}_i Y_i}{\sum_i \hat{w}_i} - \frac{\sum_i \hat{w}_i (\hat{\xi}_{ik} - x)}{\sum_i \hat{w}_i} \hat{f}'_k(x), \quad (44)$$

where

$$\hat{f}'_k(x) = \frac{\sum_i \hat{w}_i (\hat{\xi}_{ik} - x) Y_i - \{\sum_i \hat{w}_i (\hat{\xi}_{ik} - x) \sum_i \hat{w}_i Y_i\} / \sum_i \hat{w}_i}{\sum_i \hat{w}_i (\hat{\xi}_{ik} - x)^2 - \{\sum_i \hat{w}_i (\hat{\xi}_{ik} - x)\}^2 / \sum_i \hat{w}_i}. \quad (45)$$

Let $\tilde{f}_k(x)$ be a hypothetical estimator, obtained by substituting the true values w_i and ξ_{ik} for \hat{w}_i , $\hat{\xi}_{ik}$ in (44) and (45). To evaluate $|\hat{f}_k(x) - \tilde{f}_k(x)|$, one has to quantify the

orders of the differences

$$\begin{aligned} D_1 &= \sum_{\mathbf{i}} (\hat{w}_{\mathbf{i}} - w_{\mathbf{i}}), & D_2 &= \sum_{\mathbf{i}} (\hat{w}_{\mathbf{i}} - w_{\mathbf{i}}) Y_{\mathbf{i}}, \\ D_3 &= \sum_{\mathbf{i}} (\hat{w}_{\mathbf{i}} \hat{\xi}_{\mathbf{i}\mathbf{k}} - w_{\mathbf{i}} \xi_{\mathbf{i}\mathbf{k}}), & D_4 &= \sum_{\mathbf{i}} (\hat{w}_{\mathbf{i}} \hat{\xi}_{\mathbf{i}\mathbf{k}}^2 - w_{\mathbf{i}} \xi_{\mathbf{i}\mathbf{k}}^2). \end{aligned}$$

Considering D_1 , without loss of generality, assume the compact support of K_1 is $[-1, 1]$.

Since K_1 is Lipschitz continuous on its support,

$$D_1 \leq \frac{c}{nh_{\mathbf{k}}^2} \sum_{\mathbf{i}} |\hat{\xi}_{\mathbf{i}\mathbf{k}} - \xi_{\mathbf{i}\mathbf{k}}| \{I(|x - \xi_{\mathbf{i}\mathbf{k}}| \leq h_{\mathbf{k}}) + I(|x - \hat{\xi}_{\mathbf{i}\mathbf{k}}| \leq h_{\mathbf{k}})\}, \quad (46)$$

for some $c > 0$, where $I(\cdot)$ is an indicator function. Lemma 2 implies for the first term on the r.h.s. of (46)

$$\frac{1}{nh_{\mathbf{k}}^2} \sum_{\mathbf{i}} |\hat{\xi}_{\mathbf{i}\mathbf{k}} - \xi_{\mathbf{i}\mathbf{k}}| I(|x - \xi_{\mathbf{i}\mathbf{k}}| \leq h_{\mathbf{k}}) \leq \sum_{\mathbf{k}=1}^5 Z_{\mathbf{k}}^{(\cdot)} \frac{1}{nh_{\mathbf{k}}^2} \sum_{\mathbf{i}} \theta_{\mathbf{i}\mathbf{k}}^{(\cdot)} I(|x - \xi_{\mathbf{i}\mathbf{k}}| \leq h_{\mathbf{k}}).$$

Applying the central limit theorem for a random number of summands (Billingsley, 1995, page 380), observing $\sum_{\mathbf{i}} I(|x - \xi_{\mathbf{i}\mathbf{k}}| \leq h_{\mathbf{k}}) / (nh_{\mathbf{k}}) \xrightarrow{p} 2p_{\mathbf{k}}(x)$, one finds

$$\frac{1}{nh_{\mathbf{k}}} \sum_{\mathbf{i}} \theta_{\mathbf{i}\mathbf{k}}^{(\cdot)} I(|x - \xi_{\mathbf{i}\mathbf{k}}| \leq h_{\mathbf{k}}) \xrightarrow{p} 2p_{\mathbf{k}}(x) E(\theta_{\mathbf{i}\mathbf{k}}^{(\cdot)}), \quad (47)$$

provided that $E(\theta_{\mathbf{i}\mathbf{k}}^{(\cdot)}) < \infty$ for $\ell = 1, \dots, 5$. Note that $E\theta_{\mathbf{i}\mathbf{k}}^{(1)} < \infty$, $E\theta_{\mathbf{i}\mathbf{k}}^{(3)} < \infty$ by (A4), $E\theta_{\mathbf{i}\mathbf{k}}^{(4)} \leq 2\sigma_{\mathbf{X}} \sqrt{\mathbf{X}^*}$ and $E\theta_{\mathbf{i}\mathbf{k}}^{(5)} \leq |S|\sigma_{\mathbf{X}}$ by the Cauchy-Schwarz inequality. Then

$$\begin{aligned} Z_{\mathbf{k}}^{(1)} \frac{1}{nh_{\mathbf{k}}^2} \sum_{\mathbf{i}} \theta_{\mathbf{i}\mathbf{k}}^{(1)} I(|x - \xi_{\mathbf{i}\mathbf{k}}| \leq h_{\mathbf{k}}) &= O_{\mathbf{p}} \left\{ \frac{\pi_{\mathbf{k}}^{\mathbf{x}}}{\sqrt{nh_{\mathbf{k}}^2} h_{\mathbf{k}}} p_{\mathbf{k}}(x) \right\}, \\ Z_{\mathbf{k}}^{(2)} \frac{1}{nh_{\mathbf{k}}^2} \sum_{\mathbf{i}} \theta_{\mathbf{i}\mathbf{k}}^{(2)} I(|x - \xi_{\mathbf{i}\mathbf{k}}| \leq h_{\mathbf{k}}) &= O_{\mathbf{p}} \left\{ \frac{1}{\sqrt{nb_{\mathbf{X}}} h_{\mathbf{k}}} p_{\mathbf{k}}(x) \right\}, \\ Z_{\mathbf{k}}^{(3)} \frac{1}{nh_{\mathbf{k}}^2} \sum_{\mathbf{i}} \theta_{\mathbf{i}\mathbf{k}}^{(3)} I(|x - \xi_{\mathbf{i}\mathbf{k}}| \leq h_{\mathbf{k}}) &= O_{\mathbf{p}} \left\{ \frac{\|\phi_{\mathbf{k}}\|_{\infty} \mathbf{X}^*}{h_{\mathbf{k}}} p_{\mathbf{k}}(x) \right\}, \\ Z_{\mathbf{k}}^{(4)} \frac{1}{nh_{\mathbf{k}}^2} \sum_{\mathbf{i}} \theta_{\mathbf{i}\mathbf{k}}^{(4)} I(|x - \xi_{\mathbf{i}\mathbf{k}}| \leq h_{\mathbf{k}}) &= O_{\mathbf{p}} \left\{ \frac{\sqrt{\mathbf{X}^*}}{h_{\mathbf{k}}} p_{\mathbf{k}}(x) \right\}, \\ Z_{\mathbf{k}}^{(5)} \frac{1}{nh_{\mathbf{k}}^2} \sum_{\mathbf{i}} \theta_{\mathbf{i}\mathbf{k}}^{(5)} I(|x - \xi_{\mathbf{i}\mathbf{k}}| \leq h_{\mathbf{k}}) &= O_{\mathbf{p}} \left\{ \frac{\pi_{\mathbf{k}}^{\mathbf{x}}}{\sqrt{nh_{\mathbf{k}}^2} h_{\mathbf{k}}} p_{\mathbf{k}}(x) \right\}. \end{aligned} \quad (48)$$

We now obtain $(nh_k^2)^{-1} \sum_i |\hat{\xi}_{ik} - \xi_{ik}| I(|x - \xi_{ik}| \leq h_k) = O_p(\theta_k h_k^{-1})$. The asymptotic rate of the second term can be derived analogously, observing

$$\frac{1}{nh_k} \sum_i I(|x - \hat{\xi}_{ik}| \leq h_k) \leq \frac{1}{nh_k} \sum_i \{I(|x - \xi_{ik}| \leq 2h_k) + I(\sum_{l=1}^5 \theta_{ik}^{(l)} Z_k^{(l)} > h_k)\} \xrightarrow{p} 4p_k(x),$$

leading to $(nh_k^2)^{-1} \sum_i |\hat{\xi}_{ik} - \xi_{ik}| I(|x - \hat{\xi}_{ik}| \leq h_k) = O_p(\theta_k h_k^{-1})$. Then $D_1 = O_p(\theta_k h_k^{-1})$ follows.

Analogously, one shows $D_2 = O_p(\theta_k h_k^{-1})$, applying the Cauchy-Schwarz inequality for $\theta_{ik}^{(\ell)}$, $\ell = 1, 3$, and observing the independence between Y_i and $\theta_{ik}^{(\ell)}$ for $\ell = 2, 4, 5$, given the moment condition (A4). For D_3 , observe

$$D_3 = \sum_i \{(\hat{w}_i - w_i)\xi_{ik} + (\hat{w}_i - w_i)(\hat{\xi}_{ik} - \xi_{ik}) + w_i(\hat{\xi}_{ik} - \xi_{ik})\} \equiv D_{31} + D_{32} + D_{33}.$$

Then $D_{31} = O_p(\theta_k h_k^{-1})$, analogously to D_1 . It is easy to see that $D_{32} = o_p(D_{31})$. Since $D_{33} \leq c \sum_{l=1}^5 Z_k^{(l)} (nh_k)^{-1} \sum_i \theta_{ik}^{(l)} I(|x - \xi_{ik}| \leq h_k)$ for some $c > 0$, one also has $D_{33} = o_p(D_{31})$. This results in $D_3 = O_p(\theta_k h_k^{-1})$. Observing $|\hat{\xi}_{ik}^2 - \xi_{ik}^2| \leq |\hat{\xi}_{ik} - \xi_{ik}| \cdot |\xi_{ik}| + (\hat{\xi}_{ik} - \xi_{ik})^2$, one can show $D_4 = O_p(\theta_k h_k^{-1})$, using similar arguments as for D_3 , and $E\xi_{ik}^4 < \infty$ from (A4). Combining the results for D , $\ell = 1, \dots, 4$, and applying Slutsky's Theorem leads to $|\hat{f}_k(x) - \tilde{f}_k(x)| = O_p(\theta_k h_k^{-1})$. Using (A5), and applying standard asymptotic results for the hypothetical local linear smoother $\tilde{f}_k(x)$ completes the proof of (18).

To derive (19), additionally one only needs to consider $\sum_i (\hat{w}_i \hat{\zeta}_{im} - w_i \zeta_{im}) = \sum_i \{(\hat{w}_i - w_i)\zeta_{im} + (\hat{w}_i - w_i)(\hat{\zeta}_{im} - \zeta_{im}) + w_i(\hat{\zeta}_{im} - \zeta_{im})\}$, where the third term yields an extra term of order $O_p(\vartheta_{mk})$ by observing

$$|\sum_i w_i(\hat{\zeta}_{im} - \zeta_{im})| \leq \sum_{m=1}^5 Q_m^{(\cdot)} \sum_i w_i \vartheta_{im}^{(\cdot)} \leq \frac{1}{nh_{mk}} \sum_{m=1}^5 Q_m^{(\cdot)} \sum_i \vartheta_{im}^{(\cdot)} I(|x - \xi_{ik}| \leq h_{mk})$$

Similar arguments as above complete this derivation. \square

Proof of Theore

Using (A7), the derivation of θ_n^* in (42) is straightforward,

following the above arguments. To obtain (21), note that

$$\begin{aligned}
& \widehat{E}\{Y(t)|X\} - E\{Y(t)|X\} \\
\leq & \sum_{k=1}^K \sum_{m=1}^M |\widehat{f}_{mk}(\xi_k)\widehat{\psi}_m(t) - f_{mk}(\xi_k)\psi_m(t)| + \left| \sum_{k \geq K+1} \sum_{m \geq M+1} f_{mk}(\xi_k)\psi_m(t) \right| \\
\leq & \sum_{k=1}^K \sum_{m=1}^M [|\widehat{f}_{mk}(\xi_k) - f_{mk}(\xi_k)|\{|\psi_m(t)| + |\widehat{\psi}_m(t) - \psi_m(t)|\} + |f_{mk}(\xi_k)| \cdot |\widehat{\psi}_m(t) - \psi_m(t)|] \\
& + \left| \sum_{(\mathbf{k}, \mathbf{m}) \in \mathcal{N}^2 \setminus \mathcal{N}_K \times \mathcal{N}_M} f_{mk}(\xi_k)\psi_m(t) \right|.
\end{aligned}$$

This implies the convergence rate ϑ_n^* in (42).